

HW 2: Solution

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Problem 1

Since A, B are independent events

$$P(AB) = P(A)P(B) \quad \text{or} \quad \frac{|AB|}{p} = \frac{|A|}{p} \frac{|B|}{p},$$

i.e., $|AB|p = |A||B|$. Therefore, p is a divisor of $|A||B|$, assuming that $|A||B| > 0$. Since p is prime, p is a divisor of at least one of $|A|$ and $|B|$ (this elementary fact is the so-called *Euclid's Lemma* in number theory) and because $p \geq \max\{|A|, |B|\}$ the conclusion follows.

Problem 2

By the hint, it is easy to see that $P(\mathbb{1}_R(e) = 1) = \frac{1}{2}, \forall e \in E$. Therefore, $E[N_R] = \frac{|E|}{2}$. This implies that there exists at least one $R \subseteq V$ such that $N_R \geq \frac{|E|}{2}$.

Problem 3

By conditioning on the first toss we obtain

$$p_{kr} = pp_{k-1,r} + (1-p)p_{k,r-1}.$$

Additionally, $p_{k0} = 0$ and $p_{0r} = 1$ for $k, r \geq 1$.

Problem 4

By the provided hint $P(Z_i = 1) = P(R_i = 1)P(Y_i = 1) = \frac{p_1}{p_2}p_2 = p_1, \forall i$. Therefore, the Z_i 's are i.i.d. and have the same distribution with the X_i 's. This implies that $Z = (Z_1, Z_2, \dots, Z_n)$ and $X = (X_1, X_2, \dots, X_n)$ are identically distributed vectors. Moreover, by construction, $Z_i \leq Y_i, \forall i$. Therefore, $Z \leq Y$ and since f is increasing, $f(Z) \leq f(Y)$, for every realization of Z, Y . By taking expectations to both sides, we conclude that $E[f(Z)] = E[f(X)] \leq E[f(Y)]$.

Problem 5

For the first part, note that any real number x can be written as

$$x = \int_0^x 1 dt = \int_0^\infty \mathbb{1}(x > t) dt.$$

By replacing x with X and by taking the expectation to both sides the first identity follows. To exchange the order of expectation and integration we employ the *Fubini-Tonelli theorem*.

For the second part, we use the first part to write

$$E[|Y|^p] = \int_0^\infty P(|Y|^p > \tilde{t}) d\tilde{t} = \int_0^\infty P(|Y| > \sqrt[p]{\tilde{t}}) d\tilde{t}.$$

By changing variables to $t = \sqrt[p]{\tilde{t}}$ and by noting that $pt^{p-1} dt = d\tilde{t}$ the second identity follows.

Problem 6

By conditioning on any of the variables, say on X_n , the recursion in the hint follows. Clearly, $Q_0(x) = 1, \forall x \in (0, 1]$. This shows that $Q_1(x) = x, Q_2(x) = \frac{x^2}{2!}, \dots, Q_n(x) = \frac{x^n}{n!}$ (can be verified inductively). By employing the note in the previous problem, $E[N] = \sum_{n=0}^{\infty} P(N > n) = \sum_{n=0}^{\infty} \frac{x^n}{n!} = e^x$.

Problem 7

(a)

$$\begin{aligned} E[e^{\lambda x}] &= \frac{1}{2} [e^{\lambda} + e^{-\lambda}] = \frac{1}{2} \left[\sum_{k=0}^{\infty} \frac{\lambda^k}{k!} + \sum_{k=0}^{\infty} \frac{(-\lambda)^k}{k!} \right] = \sum_{k=0}^{\infty} \frac{\lambda^{2k}}{(2k)!} \\ &\leq 1 + \sum_{k=1}^{\infty} \frac{\lambda^{2k}}{2^k k!} = 1 + \sum_{k=1}^{\infty} \frac{(\lambda^2/2)^k}{k!} = e^{\frac{\lambda^2}{2}}, \end{aligned}$$

where we have used the observation that $(2k)! = k!(k+1) \cdots (2k) \geq k!2^k, \forall k \geq 1$.

(b) i.

$$E_X [e^{\lambda X}] = E_X [e^{\lambda(X - E_{X'}[X'])}] \leq E_{X, X'} [e^{\lambda(X - X')}],$$

where $E[X'] = 0$ and Jensen's inequality have been used.

ii.

$$\begin{aligned} E_{X, X'} [e^{\lambda(X - X')}] &= E_{X, X', \varepsilon} [e^{\lambda \varepsilon (X - X')}] = E_{X, X'} [E_{\varepsilon} [e^{\lambda \varepsilon (X - X')}]] \\ &\leq E_{X, X'} \left[e^{\frac{\lambda^2 (X - X')^2}{2}} \right]. \end{aligned}$$

Since $X \in [a, b]$, we have that $|X - X'| \leq (b - a)$ and therefore, $E_{X, X'} \left[e^{\frac{\lambda^2 (X - X')^2}{2}} \right] \leq E_{X, X'} \left[e^{\frac{\lambda^2 (b - a)^2}{2}} \right] = e^{\frac{\lambda^2 (b - a)^2}{2}}$. Hence, $E [e^{\lambda X}] \leq e^{\frac{\lambda^2 (b - a)^2}{2}}$.

This bound has been sharpened in class via Hoeffding's Lemma.

Problem 8

$$\begin{aligned} E[M] &= \frac{1}{n} \sum_{i=1}^n E \left[\frac{g(Y_i) f_X(Y_i)}{f_Y(Y_i)} \right] = \frac{1}{n} n E \left[\frac{g(Y) f_X(Y)}{f_Y(Y)} \right] \\ &= \int \frac{g(y) f_X(y)}{f_Y(y)} f_Y(y) dy = \int g(x) f_X(x) dx, \end{aligned}$$

where the last part is due to the equivalence of f_X, f_Y .

For the variance,

$$\begin{aligned} \text{Var}(M) &= \text{Var} \left(\frac{1}{n} \sum_{i=1}^n \frac{g(Y_i) f_X(Y_i)}{f_Y(Y_i)} \right) = \frac{1}{n^2} n \text{Var} \left(\frac{g(Y) f_X(Y)}{f_Y(Y)} \right) \\ &= \frac{1}{n} \left(E \left[\frac{g^2(Y) f_X^2(Y)}{f_Y^2(Y)} \right] - (E[g(X)])^2 \right). \end{aligned}$$

Here, we have used the fact that $\frac{g(Y_i) f_X(Y_i)}{f_Y(Y_i)}, i = 1, 2, \dots, n$ are i.i.d. random variables and the result for $E[M]$.